



## **Mathematics**

Advanced GCE

Unit 4734: Probability and Statistics 3

## Mark Scheme for January 2011

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1(i)	$Est\mu = sample mean = 5.25$	B1 1	
(ii)	Use (i) $\pm z$ SD SD = 0.19/ $\sqrt{5}$ z = 1.96 5.083 < $\mu$ < 5.417	M1 B1 B1 A1 <b>4</b> [5]	With √5 seen Rounding to 5.08, 5.42
2	Use $G - M \sim N(-6.23, \sigma^2)$ $\sigma^2 = 6.87^2 + 10.25^2$ $z = (16.23) / \sigma$ = 1.315 Probability = 0.0942 or 0.0943	M1 A1 M1 A1 A1 A1 [5]	<b>Or G-M-10~</b> N(-16.23, σ <sup>2</sup> ) Accept 0.094
3(i)	$\int_{0}^{2} a e^{-1} dt + \int_{2}^{\infty} a e^{-\frac{1}{2}t} dt = 1$ [ $a e^{-1}t$ ] + [-2 $a e^{-\frac{1}{2}t}$ ] => $a = \frac{1}{4}e$ AG	M1 A1 A1 <b>3</b>	Properly obtained
(ii)	$\int_{q_3}^{\infty} \frac{1}{4} e^{1 - \frac{1}{2}t} dt = \frac{1}{4}$ $\begin{bmatrix} -\frac{1}{2}e^{1 - \frac{1}{2}t} \end{bmatrix}$ $\frac{1}{2}q_3 + 1 = -\ln 2$ $\Rightarrow q_3 = 2(\ln 2 + 1) \text{ or } 3.39$	M1 B1 M1 A1 <b>4</b> [7]	OR $\int_{0}^{2} \frac{1}{4} dt + \int_{2}^{q} \frac{1}{4} e^{1-t/2} dt = \frac{3}{4}$ AEF For taking logs (not ln(- )) AEF
4	$\hat{p}_2 = 106/143, \ \hat{p}_1 = 61/107$ = 0.7413 = 0.5701 Pooled est $p = 167/250$ Variance est = $\binom{167}{250}\binom{83}{250}(143^{-1}+107^{-1})$ Test statistic $z = (0.7413 - 0.5701)/SD$ = 2.84(35) Smallest significance level = 0.23% SR: No pe, B1B0B0M1A1(2.84)M1A1 Max 5/7	B1 B1 B1 M1 A1 M1 A1√ [7]	For both Only if used ART 0.22 or 0.23 Accept 0.0023 $\sqrt{z}$ M1A0 if 0.25%
5(i)	$s^2 = 0.2 \times 0.8/90$ $p_s \pm zs$ z = 1.645 $0.1306 < p_y < 0.2693$	B1 M1 B1 A1 <b>4</b>	OR /89 Art (0.131, 0.269)
(ii)	0.7306 < <i>p</i> <sub>p</sub> < 0.8694	B1ft <b>1</b>	ft (i) Art (0.731, 0 <b>.869</b> )
(iii)	If a large number of such intervals were calculated from independent samples, approximately 90% of all such intervals would contain <i>p</i>	B2 2	Or: Probability that such an interval contains <i>p</i> is 0.9 B1 for right idea
(iv)	(0.131, 0.269) encloses 0.25 so Mendel's theory is supported	M1 A1 √ 2 [9]	Or equivalent Ft CI(i)

6(i)	$\mathbf{G}(\mathbf{y}) = \mathbf{P}(\mathbf{Y} \le \mathbf{y})$	M1	
0(1)	$G(y) = P(X \ge 1/y)$ $= P(X \ge 1/y)$	A1	
	= 1 - F(1/y) = 1 - F(1/y)	M1	
	$= 1 - \Gamma(1/y)$ = $(2y - 1)/(y+1)$	Al	
	-(2y-1)/(y+1) For $\frac{1}{2} \le \frac{1}{y} \le 2 = \frac{1}{2} \le y \le 2$	B1	Seen
	X and Y have identical distributions	B1 6	Seen
	X and T have identical distributions	DI U	
	SR: CDF not used.		
	y decreases with x		
	Use $g(y) = f(x(y) dx/dy) $	M1	
	$f(x) = 3/(x+1)^2$	M1A1	
	$ dx/dy =1/y^2$	B1	
	$g(y) = [3/(y^{-1}+1)^2][1/y^2] = 3/(y+1)^2$ ; for $\frac{1}{2} \le y \le 2$	M1A1B1	
	So X and Y have identical distributions	B1 8	
	$\gamma = 2 \gamma + $		
(ii)	$f(x) = F'(x) = 3/(x+1)^2, \frac{1}{2} \le x \le 2$	M1A1	Must have range of $x$
	$E(X+1) = \int_{1}^{2} \frac{3}{1-x^2} dx$	N/1	AEF Not if awarded in (i)
	$E(X+1) = \int_{\frac{1}{2}}^{2} \frac{3}{x+1} dx$	M1	
	$=3\ln 2$ (2.08)	A1	
		AI	
	E(1/X) = E(X)	M1	
	$= 3\ln 2 - 1 (1.08)$	A1 6	
		[12]	
7(i)	In a 2×2 contingency table	B1 1	Or equivalent Accept df=1
(ii)	H <sub>0</sub> : Vaccine type and outcome are independent	B1M*dep	Accept omission of H <sub>1</sub>
	H <sub>1</sub> : They are not independent		
	E-values: 10.81 12.19	M1	1 correct E value
	318.19 358.81	A1	Accept 1 dp
	$\chi^2 = 7.69^{\ 2}(10.81^{-1} + 12.19^{-1} + 318.19^{-1} + 358.81^{-1})$	M1	1 correct $\chi^2$ value ft E values
	10.77	M1	Using Yates' correctly
	=10.67	A1	Accept 10.7
	CV = 6.635	B1	
	<b>10.67</b> > CV	M1	
	Reject $H_0$ , there is sufficient evidence at the 1%		
	significance level that the outcome of the test depends	A 1./	$\sqrt{10.67}$
	on the vaccine used	A1√	V 10.07
	$T_{1}$ = 1.1 $(1 - 1)$ $(1 - 1)$	dep*M	
	The results is significant at a level less than $\frac{1}{2}$ %, so	A1./ 10	G
	the evidence is very strong	A1 √ <b>10</b>	Sensible comment. $\sqrt{10.67}$
		[11]	

8(i)	When independent samples are drawn from populations having a common variance	B1 1	For common variance
(ii)	(a) Lung capacities should have normal distributions with a common variance	B1 1	Normal distributions required In context here
	<b>(b)</b> $H_0:\mu_1=\mu_2$ , $H_1:\mu_1 > \mu_2$	B1	Or equivalent
	$s_1^2 = \frac{1}{19}(90.43 - 42.4^2 / 20)$	M1	For 1 correct $s^2$
	$s_2^2 = \frac{1}{21}(82.93 - 42.5^2 / 22)$ $\overline{x}_1 = 2.12 \qquad \overline{x}_2 = 1.93(2)$ PEV, $s^2 = (19s_1^2 + 21s_2^2)/(20 + 22 - 2)$ = 0.03424(3)	B1 M1 A1	For both
	Test statistic = $\frac{2.12 \cdot 1.932}{\sqrt{s^2 (20^{-1} + 22^{-1})}}$	M1A1	ft s <sup>2</sup>
	= 3.29(15) CV = 2.423 TS > CV	A1√ B1 M1	Accept answer rounding to 3.3 If z used the B0M0A0 Compare with CV
	There is sufficient evidence at the 1% SL that the mean lung capacity is greater for children whose parents do not smoke than for children whose parents do smoke SR1: For 2-tail test Lose 1 <sup>st</sup> B1 and last 3. Max 8/11 SR2: If $s^2 = s_1^2/20 + s_2^2/22$ , B1M1A0A0M1A0A1(3.32) B1M1A1 Max 8/11	A1 11	Or equivalent, in context
	(c) $t = 2.704$ 0.1882 $\pm ts(20^{-1}+22^{-1})^{1/2}$	B1 M1	Accept 0.19
	(0.0336, 0.3423)	A1 3 [16]	(0.033-0.036,0.342-0.346)

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